

The Benefits of Probability Proportional to Size Sampling in Cluster-Randomized Experiments

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Aug 29, 2014

Cluster-Randomized Experiments

- Cluster-randomized experiments: Randomization of treatment across clusters of units instead of the units themselves.
- Reasons for cluster-level randomization: Logistical, not to improve power.

E.g.: Impractical or infeasible to randomize across units, avoid “interference” or “treatment contamination,” etc.

Cluster-Randomized Experiments: Examples

Examples:

- Assign treatment to villages, interested in effects at individual/household level.

National Solidarity Programme (Beath, Christia, Enikolopov)
Clientelism and Vote Buying (Wanchekon)

- Assign treatment to households, interested in individual effects.
Get out the vote (Gerber, Green, Larimer)

Also common in medical trials and experiments in education.

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- Current practice: **clusters sampled using SRS**
- Horvitz-Thompson: unbiased, but not location invariant.
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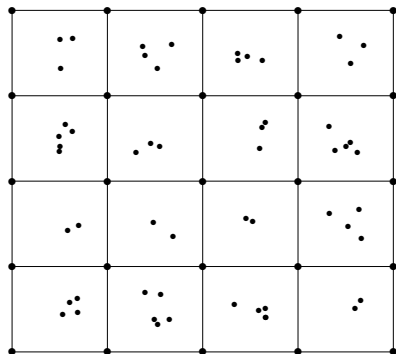
Difference-in-means: Location invariant but biased when treatment effects correlated with number of units within cluster (Middleton and Aronow).
- **Solution:** Sampling clusters with probability proportional to size—the number of units within each clusters—eliminate these problems with estimation.

Cluster-Randomized Experiments: Framework

- The population of units of interest y_1, \dots, y_n partitioned into $\#c$ clusters.
- Quantity of interest: population average treatment effect (PATE).

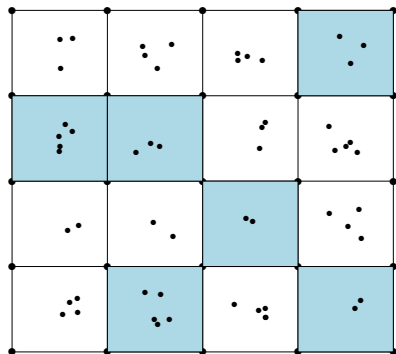
$$PATE = \frac{1}{n} \sum_{i=1}^n y_i(1) - \frac{1}{n} \sum_{i=1}^n y_i(0)$$

$y_i(1)$ is response of unit i under treatment, $y_i(0)$ is response under control.



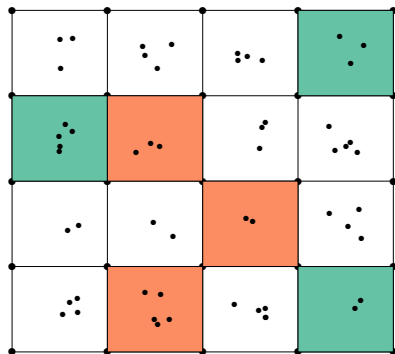
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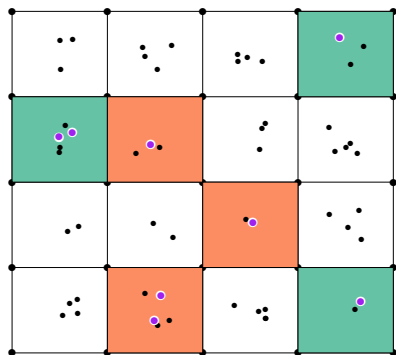
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- 2 Randomize treatment across clusters.
- 3 Sample $\#s_c$ units within each of the sampled clusters.

Inference on PATE made using purple sampled units.



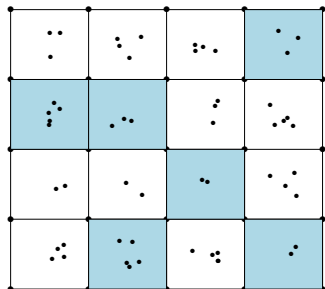
Current practice: SRS sampling of clusters

- Under Neyman-Rubin potential outcomes model:

$$Y_i = y_i(1)T_i + y_i(0)(1 - T_i)$$

Under SRS of clusters, no current estimator of PATE that is both unbiased and invariant to location shifts of potential outcomes...

- ... without the introduction of additional parameters (Des Raj) (Middleton and Aronow).



Inference on PATE: Horvitz-Thompson

- Horvitz-Thompson estimator of PATE:

$$\sum_{c \in \#s} w_c \left(\sum_{T_c=1} \frac{1}{\#T_1} \frac{n_c}{n} \sum_{k \in c} \frac{y_{kc}(1)}{\#s_c} - \sum_{T_c=0} \frac{1}{\#T_0} \frac{n_c}{n} \sum_{k \in c} \frac{y_{kc}(0)}{\#s_c} \right)$$

$$\sum_{k \in c} \frac{y_{kc}(1)}{\#s_c} = \text{Within-cluster sample mean}$$

$$\frac{n_c}{n} = \text{Fraction of units contained in cluster } c$$

$$\frac{1}{\#T_1} = \text{Number of treated clusters}$$

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- Unbiased, but not invariant to location shifts of potential outcomes, which inflates variance:

$$y_{kc}^* = y_{kc} + \alpha.$$

$$\widehat{PATE}^* = \widehat{PATE} + k\alpha (\#n_1 - \#n_0).$$

- E.g.: Adjustments for inflation or currency for logged responses.
Recoding binary variable from 0–1 to 1–0.

Inference on PATE: Difference-in-means

- Difference-in-means estimator of PATE:

$$\frac{\sum_{T_c=1} \sum_{k \in C} y_{kc}(1)}{\#s_1} - \frac{\sum_{T_c=0} \sum_{k \in C} y_{kc}(0)}{\#s_0}$$

- Invariant to location shifts, but biased if treatment effects are correlated with cluster sizes.

E.g. Treatment effect smaller going from small towns to large cities.

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- Common to block or stratify on cluster size to reduce bias.
- Suboptimal if other covariates predict treatment effects better.
- Still some bias in finite experiments.

Solution: Better design

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- Focus attention on sampling without replacement (PPSWOR sampling):
Fixed number of clusters sampled \implies Easier to design experiment with budget constraint.

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- Does not uniquely define a sample; joint probabilities $\pi_{cc'}$ need to be specified.
- `SunterSampling` R package; efficiently draw approximate PPSWOR sample with nice properties.

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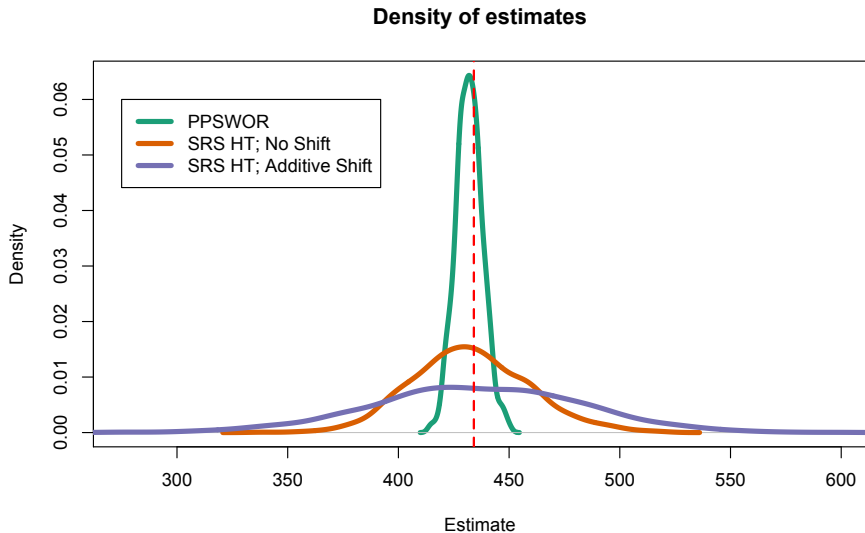
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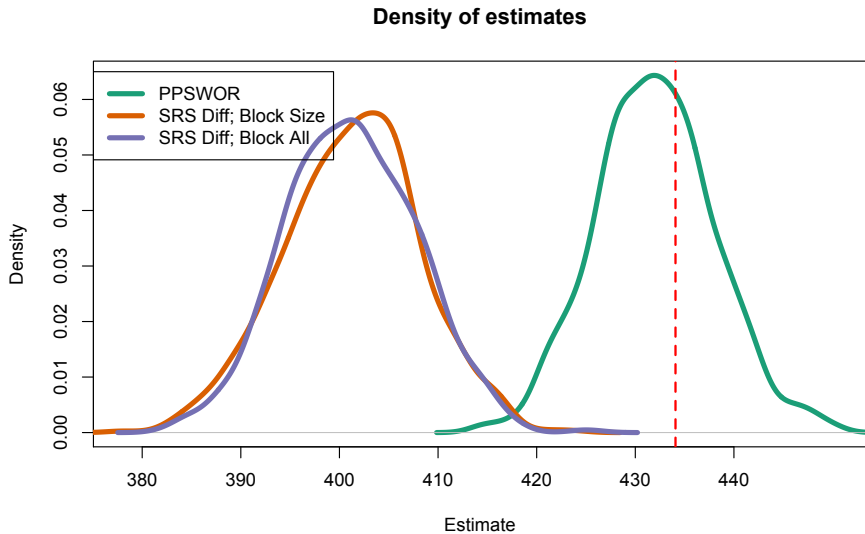
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- Unbiased and invariant to location shifts of potential outcomes.
- Freedom to block on covariates other than cluster size.
- Variance not dramatically different from SRS methods.
- Equivalent to difference-in-means estimator when same within-cluster sample sizes or under matched-pairs blocking.

Simulation results



Simulation results



Extensions

Results hold when:

- Stratifying clusters before sampling.
E.g.: Sample large cities and small villages separately.
- Stratifying within-cluster sample.
E.g.: Gender, race.
- When treatment is randomized across individual units.
Useful when design has treatments randomized at both cluster and treatment level.

Thank you.